

Assignment 2 Outline Solution

A2 – 1. This question involves a $k = 3$ variable multinomial distribution as a probability model.

$$(a) \Pr(3G, 2R, 1U) = \binom{6}{3 \ 2 \ 1} \left(\frac{1}{2}\right)^3 \left(\frac{1}{3}\right)^2 \left(\frac{1}{6}\right)^1 = \frac{5}{36} = 0.138\bar{8}.$$

$$(b) \Pr(\text{at least one } R \text{ and exactly one } U) =$$

$$\begin{aligned} & \Pr(4G, 1R, 1U) + \Pr(3G, 2R, 1U) + \Pr(2G, 3R, 1U) + \Pr(1G, 4R, 1U) + \Pr(0G, 5R, 1U) \\ &= \frac{5}{48} + \frac{5}{36} + \frac{5}{54} + \frac{5}{162} + \frac{1}{243} = \frac{2,882}{7,776} \approx 0.370\ 628. \end{aligned}$$

$$\text{Alternatively: } \Pr(\text{at least one } R \text{ and exactly one } U) =$$

$$= \Pr(1U) - \Pr(5G, 0R, 1U) = \left(\frac{5}{6}\right)^5 - \left(\frac{1}{2}\right)^5 = \frac{5^5 - 3^5}{6^5} = \frac{2,882}{7,776} \quad (\text{as before}).$$

$$(c) \Pr(\text{at least one } R) = 1 - \Pr(\text{no } R) = 1 - \left(\frac{2}{3}\right)^6 = \frac{665}{729} \approx 0.912\ 209.$$

(d) To be sure of satisfying the conditions of the event whose probability is required here, we can list all 28 of the relevant possibilities for a total of 6 items, as follows:

G	R	U	G	R	U	G	R	U	G	R	U	G	R	U	G	R	U
6	0	0	5	1	0	4	2	0	4	1	1	3	2	1	3	3	0
0	6	0	1	5	0	2	4	0	1	4	1	2	3	1	3	0	3
0	0	6	0	5	1	0	4	2	1	1	4	1	3	2	0	3	3
			0	1	5	0	2	4				1	2	3			
			5	0	1	4	0	2				3	1	2	2	2	2
			1	0	5	2	0	4				2	1	3			

$$\text{Hence: } \Pr(\text{at least one item which is either } R \text{ or } U) =$$

$$= \Pr[(\text{at least one } R) \cup (\text{at least one } U)] = 1 - \Pr(6G) = 1 - \left(\frac{1}{2}\right)^6 = \frac{63}{64} = 0.984\ 375.$$

A2 – 2. This question also involves the multinomial distribution as a probability model; there is an *implied* (fourth) category of 6-9 blackflies in a tent, with a frequency of 1. The probabilities of the $k = 5$ outcomes of this multinomial distribution are obtained from an appropriate *Poisson* model.

With an average of 15 blackflies per cubic metre of air, we expect 9 blackflies in a tent of volume 0.6 m^3

$$\text{Hence: } \Pr(y \text{ blackflies in a tent of volume } 0.6 \text{ m}^3) = \frac{9^y e^{-9}}{y!}; \quad y = 0, 1, 2, 3, \dots$$

Thus, the probabilities of the five outcomes for the multinomial distribution are:

$$\Pr(0 \text{ blackflies in a tent}) = e^{-9} \approx 0.000\ 123\ 410 = \pi_0, \text{ say;}$$

$$\Pr(1 \text{ or } 2 \text{ blackflies in a tent}) = \Pr(1 \text{ blackfly}) + \Pr(2 \text{ blackflies})$$

$$= 9e^{-9}/1! + 9^2e^{-9}/2! = \frac{99}{2}e^{-9} \approx 0.006\ 108\ 785 = \pi_2, \text{ say;}$$

$$\Pr(3, 4 \text{ or } 5 \text{ blackflies in a tent}) = \Pr(3 \text{ blackflies}) + \Pr(4 \text{ blackflies}) + \Pr(5 \text{ blackflies})$$

$$= 9^3e^{-9}/3! + 9^4e^{-9}/4! + 9^5e^{-9}/5! = \frac{17,739}{20}e^{-9} \approx 0.109\ 458\ 325 = \pi_3, \text{ say;}$$

$$\Pr(6, 7, 8 \text{ or } 9 \text{ blackflies in a tent}) = \Pr(6 \text{ blackflies}) + \Pr(7 \text{ blackflies}) + \Pr(8 \text{ blackflies}) + \Pr(9 \text{ blackflies})$$

$$= 9^6e^{-9}/6! + 9^7e^{-9}/7! + 9^8e^{-9}/8! + 9^9e^{-9}/9! = \frac{1,712,421}{448}e^{-9} \approx 0.471\ 717\ 723 = \pi_4, \text{ say;}$$

$$\Pr(10 \text{ or more blackflies in a tent}) = 1 - \pi_0 - \pi_2 - \pi_3 - \pi_4 \approx 0.412\ 591\ 755 = \pi_5, \text{ say.}$$

(a) Then, with the categories in the same order as in the question and in the preceding Poisson calculations, and using the usual notation for the multinomial distribution:

$$\Pr(6, 4, 3, 1, 6) = \binom{20}{6 \ 4 \ 3 \ 1 \ 6} \pi_0^6 \pi_2^4 \pi_3^3 \pi_4^1 \pi_5^6 \approx 4.892\ 901 \times 10^{-28}.$$

(b) Conditioning on six tents having no blackflies as event B in the definition of $\Pr(A|B)$, we have:

$$\begin{aligned} \Pr(4, 3, 1, 6|6) &= \frac{\Pr(6, 4, 3, 1, 6)}{\Pr(\text{no blackflies in 6 tents})} = \frac{\Pr(6, 4, 3, 1, 6)}{\binom{20}{6} \pi_0^6 (1 - \pi_0)^{14}} \\ &\approx \frac{4.892\ 900\ 986 \times 10^{-28}}{1.366\ 883\ 032 \times 10^{-19}} \approx 3.579\ 605 \times 10^{-9}. \end{aligned}$$

(continued overleaf)

Assignment 2 Outline Solution (continued 1)

A2 – 3. This question involves the using the Central Limit Theorem to justify an assumed approximate normality for the distribution of a sum of 200 discrete random variables.

Let the random variable Y_i represent the number of typing mistakes on the i th page; we are given in the question that the probability function of Y_i is as shown at the right.

y_i	0	1	2	3	4	5	6
$f(y_i)$	0.05	0.25	0.31	0.22	0.10	0.05	0.02

(a) Then: $E(Y_i) = 0 \times 0.05 + 1 \times 0.25 + \dots + 6 \times 0.02 = 2.3$ mistakes (as required);

$$E(Y_i^2) = 0^2 \times 0.05 + 1^2 \times 0.25 + \dots + 6^2 \times 0.02 = 7.04 \text{ mistakes}^2$$

so that: $s.d.(Y_i) = \sqrt{7.04 - (2.3)^2} = \sqrt{1.75} \approx 1.3229$ mistakes (as required).

If the manuscript contains 200 typed pages, the total number of mistakes is given by: $T = \sum_{i=1}^{200} Y_i$,

so that: $E(T) = 460$ mistakes, $s.d.(T) = \sqrt{350} \approx 18.7083$ mistakes;

also, T has approximately a normal distribution as a consequence of the Central Limit Theorem.

(b) We want to find: $\Pr(T \geq 500) \approx \Pr[N(460, \sqrt{350}) \geq 500 - \frac{1}{2}] \approx \Pr[N(0, 1) \geq \frac{499.5 - 460}{\sqrt{350}}]$
 $\approx \Pr[N(0, 1) \geq 2.111364] \approx 0.017371$;

i.e., the probability of at least 500 mistakes in the manuscript is about 0.017 or about 1.7%.

A2 – 4. This question involves the Poisson distribution (and its normal approximation) as a probability model.

Let the random variable Y represent the number of calls arriving at the exchange in a time interval of length t ; we use the model $Y \sim \text{Po}(\lambda t)$.

i.e., $\Pr(Y \text{ calls reach the exchange in time } t) = \frac{(\lambda t)^y e^{-\lambda t}}{y!}; \quad y = 0, 1, 2, 3, \dots,$

where: λ = average rate at which calls reach the exchange
 $= 225$ per hour $= 3.75$ per minute.

We know that the Poisson distribution has mean $\mu = \lambda t = 225$ calls and standard deviation $\sqrt{\mu} = 15$ calls.

Also, to simplify the evaluation of the required Poisson probabilities, we use the normal approximation to the Poisson distribution; with a mean as large as 225, we expect a very good approximation.

(a) Then: $\Pr(Y \geq 250) \approx \Pr[N(225, 15) \geq 250 - \frac{1}{2}] \approx \Pr[N(0, 1) \geq \frac{249.5 - 225}{15}]$
 $\approx \Pr[N(0, 1) \geq 1.63] \approx 0.051199$;

i.e., the probability at least 250 calls arrive at the exchange in 1 hour is about 0.051 or about 5.1%.

(b) Also: $\Pr(Y = 250) \approx \Pr[250 - \frac{1}{2} \leq N(225, 15) \leq 250 + \frac{1}{2}] \approx \Pr[\frac{249.5 - 225}{15} \leq N(0, 1) \leq \frac{250.5 - 225}{15}]$
 $\approx \Pr[1.63 \leq N(0, 1) \leq 1.7] \approx 0.006634$;

i.e., the probability exactly 250 calls arrive at the exchange in 1 hour is about 0.0066 or about 0.7%.

NOTE: It is of interest to compare the result obtained relatively easily by means of the normal approximation in (b) to the more troublesome calculation of the exact Poisson probability:

$$\Pr(\text{exactly 250 calls in 1 hour}) = \frac{(225)^{250} e^{-225}}{250!} \approx \frac{2.134868355 \times 10^{490}}{3.232860761 \times 10^{492}} \approx 0.006603651$$

i.e., the normal approximation is accurate to about 1 part in 200 or about 0.5% in this situation.

A2 – 5. Let the random variable Y_i represent the number of people in the i th car crossing the bridge; we are given that the probability function of Y_i is as shown at the right.

y_i	1	2	3	4	5	6
$f(y_i)$	0.05	0.43	0.27	0.12	0.09	0.04

Then: $E(Y_i) = 1 \times 0.05 + 2 \times 0.43 + \dots + 6 \times 0.04 = 2.89$ people per car;

so that: $E(Y_i - 1) = 2.89 - 1 = 1.89$ passengers per car.

Hence, the following amounts are raised by the three options given:

- at \$1.00 per car, raise \$1.00 per vehicle;
- at \$0.40 per person, raise $\$0.40 \times 2.89 = \1.156 per vehicle;
- at \$0.75 per car and driver and \$0.15 per passenger, raise $\$0.75 + \$0.15 \times 1.89 = \$1.0335$ per vehicle.

Thus, the second option raises the most money.

(continued)

Assignment 2 Outline Solution (continued 2)

- A2 – 6.** (a) Let the random variable Y represent the number of blood tests required for a group of k people; Y has two possible values: 1 and $k + 1$.

Also: $\Pr(Y=1) = \Pr(\text{pooled test is negative})$
 $= \Pr(\text{none of the } k \text{ people is infected}) = (1 - \pi)^k$
 (assuming that the results of tests on different people in the group are independent);

$$\Pr(Y=k+1) = 1 - \Pr(Y=1) = 1 - (1 - \pi)^k$$

Thus, the probability function of Y is as shown at the right,

so that: $E(Y) = \sum_y y \cdot f(y) = 1 \times (1 - \pi)^k + (k + 1) \times [1 - (1 - \pi)^k]$
 $= (k + 1) - k(1 - \pi)^k$

y_i	1	$k + 1$
$f(y_i)$	$(1 - \pi)^k$	$1 - (1 - \pi)^k$

which is the expression given for the average number of tests required for a group of k people.

- (b) $E(Y)$ is the expected number of tests which will have to be made for a group of k people; hence, the number expected for n/k groups of k people is:

$$\frac{n}{k} \times E(Y) = \frac{n}{k} [(k + 1) - k(1 - \pi)^k] \quad [\text{using (a)}].$$

- (c) If π is small, $(1 - \pi)^k \approx 1 - k\pi$ (using the binomial theorem) so that, from the result in (b), the expected number of tests is given by:

$$T \approx \frac{n}{k} [(k + 1) - k(1 - k\pi)] = \frac{n}{k} [k^2\pi + 1] = n(k\pi + k^{-1}), \text{ as required.}$$

Then: $\frac{dT}{dk} \approx n\pi - \frac{n}{k^2}$, which is zero when: $k^2 = \frac{1}{\pi}$ or: $k \approx 1/\sqrt{\pi}$, as required;

also: $\frac{d^2T}{dk^2} \approx \frac{2n}{k^3}$ which is positive because n and k are positive, as required for a *minimum*.

Illustration: If $\pi = 0.01$, we take $k = 1/\sqrt{0.01} = 10$, so $T \approx n \times 10 \times 0.01 + n/10 = 0.2n$;

i.e., on average, only about *one-fifth* as many tests are needed in the grouped procedure as would be needed if n people were tested individually, when π is 0.01 or 1%.

- A2 – 7.** The random variable Y is the number of prizes awarded. To tabulate its probability function, we need to find the values Y can take on and the probabilities of these values; we can set out the relevant information as shown in the table at the right.

Using this information and remembering that there are 10,000 tickets in the lottery, we can tabulate the probability function of Y as follows:

TYPE OF TICKET NUMBER	NUMBER OF TICKETS	PRIZES
4 digits the same	$\binom{10}{1} \frac{4!}{4!} = 10$	1
3 digits the same	$\binom{10}{1} \binom{9}{1} \frac{4!}{3!1!} = 360$	4
2 pairs of digits the same	$\binom{10}{2} \frac{4!}{2!2!} = 270$	6
2 digits the same	$\binom{10}{1} \binom{9}{2} \frac{4!}{2!1!1!} = 4,320$	12
all digits different	$\binom{10}{4} \frac{4!}{1!1!1!1!} = 5,040$	24

$\frac{y}{f(y)} \quad \begin{array}{c|ccccc} & 1 & 4 & 6 & 12 & 24 \\ \hline & 0.001 & 0.036 & 0.027 & 0.432 & 0.504 \end{array}$, so that: $E(Y) = 1 \times 0.001 + 4 \times 0.036 + \dots + 24 \times 0.504 = 17.587$.

With tickets costing \$1 and each prize being \$500, the lottery operator's profit is: $W = \$10,000 - \$500 \times Y$

$$\therefore E(W) = \$10,000 - \$500 \times E(Y) = \$10,000 - \$500 \times 17.587 = \$1,206.50,$$

which is the required lottery operator's expected profit.

- A2 – 8.** This question involves the use of the Central Limit Theorem to justify an assumed approximate normality for the distribution of a sum of 400 discrete random variables.

Let the random variable Y_i represent the number of slices of pickle on the i th hamburger; from the information given in the statement of the question, we can obtain the probability function of Y_i as shown at the right.

y_i	0	1	2	3
$f(y_i)$	0.2	0.4	0.3	0.1

Then: $E(Y_i) = 0 \times 0.2 + 1 \times 0.4 + \dots + 3 \times 0.1 = 1.3$ slices;

$$E(Y_i^2) = 0^2 \times 0.2 + 1^2 \times 0.4 + \dots + 3^2 \times 0.1 = 2.5 \text{ slices}^2;$$

so that: $s.d.(Y_i) = \sqrt{2.5 - (1.3)^2} = 0.9$ slices;

(continued overleaf)

Assignment 2 Outline Solution (continued 3)

A2 – 8. If 400 hamburgers are to be sold, the total number of slices of pickle is given by: $T = \sum_{i=1}^{400} Y_i$,
(cont.) so that: $E(T) = 520$ slices, $\text{s.d.}(T) = \sqrt{324} = 18$ slices;
 also, T has approximately a normal distribution as a consequence of the Central Limit Theorem.

We want to find the value of t such that: $\Pr(T \leq t) = 0.99$ or: $\Pr(T > t) = 0.01$.

Hence: $\Pr(T > t) \approx \Pr[N(520, 18) > t + \frac{1}{2}] \approx \Pr[N(0, 1) > \frac{t + \frac{1}{2} - 520}{18}] \approx 0.01$;

$\therefore \frac{t - 519.5}{18} \approx 2.3263$ [from the $N(0, 1)$ table] so that: $t \approx 561.37 \approx 562$ (slices).

i.e., about 562 slices of pickle need to be on hand to be 99% sure of meeting the demand.

***A2 – 9.** This question involves finding the mean of a negative binomial (really a *geometric*) distribution, for which the parameter π is obtained from an appropriate *Poisson* model.

Let the random variable Y represent the number of mistakes on an equiprobably-selected page;

we use the model $Y \sim \text{Po}(2)$, so that: $\Pr(y \text{ mistakes on the page}) = \frac{2^y e^{-2}}{y!}$; $y = 0, 1, 2, 3, \dots$,

Hence, for an equiprobably-selected page: $\Pr(\text{page is acceptable}) = \Pr(\text{page has 2 or fewer mistakes})$
 $= \Pr(\text{no mistakes}) + \Pr(1 \text{ mistake}) + \Pr(2 \text{ mistakes})$
 $= 2^0 e^{-2}/0! + 2^1 e^{-2}/1! + 2^2 e^{-2}/2!$
 $= 5e^{-2} \approx 0.676\ 676 = \pi$ say;

$\Pr(\text{page is unacceptable}) = 1 - \Pr(\text{page is acceptable})$
 $= 1 - 5e^{-2} = 1 - \pi$

Then, if the random variable R represents the number of times a page must be typed before it is acceptable:

$\Pr(\text{type a page } r \text{ times}) \equiv f(r) = (1 - \pi)^{r-1} \times \pi$; $r = 1, 2, 3, \dots$

Hence: $E(R) = \sum_r r \cdot f(r) = \sum_{r=1}^{\infty} r \cdot (1 - \pi)^{r-1} \times \pi = \pi \sum_{r=1}^{\infty} r \cdot (1 - \pi)^{r-1}$
 $= \pi \frac{d}{d\pi} \sum_{r=1}^{\infty} (1 - \pi)^r$
 $= \pi \frac{d}{d\pi} \left(\frac{\pi - 1}{\pi} \right)$
 $= \frac{1}{\pi} = e^2/5 \approx 1.477\ 811$;

thus, the secretary can expect to type $100 \times 1.477\ 811 \approx 148$ pages.

NOTE: Three points of interest in the final calculation of $E(R)$ are:

- recognition that $r \cdot (1 - \pi)^{r-1}$ is the derivative with respect to π of $(1 - \pi)^r$.
- interchange of the order of summation and differentiation, which is permitted because the two operations involve *different* variables (r and π respectively);
- the result for the sum of an infinite geometric series.