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## Minimizing CVaR and VaR for a portfolio of derivatives

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### Abstract

Value at risk (VaR) and conditional value at risk (CVaR) are frequently used as risk measures in risk management. Compared to VaR, CVaR is attractive since it is a coherent risk measure. We analyze the problem of computing the optimal VaR and CVaR portfolios. We illustrate that VaR and CVaR minimization problems for *derivatives portfolios* are typically ill-posed. We propose to include cost as an additional preference criterion for the CVaR optimization problem. We demonstrate that, with the addition of a proportional cost, it is possible to compute an optimal CVaR derivative investment portfolio with significantly fewer instruments and comparable CVaR and VaR. A computational method based on a smoothing technique is proposed to solve a simulation based CVaR optimization problem efficiently. Comparison is made with the linear programming approach for solving the simulation based CVaR optimization problem.

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## 1. Introduction

One of the main objectives of risk management is to evaluate and improve the performance of financial organizations in light of the risks taken to achieve profits. A current standard benchmark for firm-wide measures of risk is value at risk (VaR) (Duffie and Pan, 1997). For a given time horizon  $\bar{t}$  and confidence level  $\beta$ , the value at risk of a portfolio is the loss in the portfolio's market value over the time horizon  $\bar{t}$  that is exceeded with probability  $1 - \beta$ . However, as a risk measure, VaR has recognized limitations. Firstly it lacks subadditivity and convexity (Artzner et al., 1997; Artzner et al., 1999). For example, the VaR of the combination of two portfolios can be greater than the sum of VaR of the individual portfolios. Indeed, VaR is a coherent risk measure only when it is based on the standard deviation of normal distributions. In addition, it has been shown in Mausser and Rosen (1999) and McKay and Keefer (1996) that the problem of minimizing VaR of a portfolio can have multiple local minimizers.

An alternative risk measure to VaR is *conditional value at risk* (CVaR), which is also known as *mean excess loss*, *mean shortfall* or *tail VaR*. For a given time horizon  $\bar{t}$  and confidence level  $\beta$ , CVaR is the conditional expectation of the loss above VaR for the time horizon  $\bar{t}$  and the confidence level  $\beta$ . The CVaR risk measure, with a slight modification, is also applicable to distributions with jumps (Rockafellar and Uryasev, 2002). It has been shown (Pflug, 2000) that CVaR is a coherent risk measure that has many attractive properties including convexity, e.g., see Ogryczak and Ruszczyński (2002) for an overview of CVaR. In addition, minimizing CVaR typically leads to a portfolio with a small VaR.

A convex optimization problem has been proposed in Rockafellar and Uryasev (2000) to compute the optimal CVaR portfolio. We describe the mathematical formulation of CVaR optimization problem in Section 2. In particular, when this optimization problem is approximated by Monte Carlo simulation, it has an equivalent linear programming formulation and can be solved using standard linear programming methods.

Derivative contracts have become increasingly important as investment tools for achieving higher returns and decreasing funding costs. In this paper, we first analyze in Section 3 the well-posedness of the optimal CVaR/VaR portfolio selection problem when the investment universe consists of derivative contracts. We illustrate that the CVaR/VaR optimization problem for derivative portfolios typically has an infinite number of solutions if the derivative values are computed using delta–gamma approximations. Moreover, even when the derivative values are computed with more accurate methods such as analytic formulae, numerical partial differential equations, or Monte Carlo methods, the CVaR/VaR optimization problem for derivative portfolios remains ill-posed in the sense that there are many portfolios that have similar CVaR/VaR values to that of the optimal portfolio and slight perturbation of the data can lead to significantly different optimal solutions. We illustrate this with derivative CVaR portfolio examples in Section 3. In Section 4 we focus on the CVaR optimization problem and introduce cost as an additional preference; the cost is modeled as proportional to the magnitude of the holding positions. A similar

consideration can be applied to the VaR optimization problem although VaR optimization is a more computationally challenging task. We show that a convex programming problem can be formulated for CVaR optimization under the proposed proportional cost model. In addition, we demonstrate that the proposed CVaR optimization formulation with cost is able to limit both the transaction cost and management cost; an optimal CVaR derivative investment portfolio using a suitable weighted cost parameter has smaller total trading positions, significantly fewer instruments, and comparable CVaR (and VaR).

The standard method for a CVaR optimization problem is a linear programming (LP) approach. Using Monte Carlo simulation, a piecewise linear function is used to approximate the typically continuous differentiable CVaR function which results in a linear programming problem. This LP is then solved using standard linear programming software. We illustrate that this approach becomes inefficient for large scale CVaR optimization problems.

A computational method based on a smoothing technique is proposed in Section 5 to efficiently solve a simulation based CVaR optimization problem. Comparison is made with the linear programming approach to solve the simulation based CVaR optimization problem. We demonstrate that the smoothing formulation, compared with the linear programming approach, is computationally much more efficient in both CPU usage and memory requirement and is capable of solving larger problems.

## 2. Mathematical formulation

For a time horizon  $\bar{t}$ , let  $f(x, S)$  denote the loss of a portfolio with decision variable  $x \in \mathbb{R}^n$  and random variable  $S \in \mathbb{R}^d$  denote the value of underlying risk factors at  $\bar{t}$ . Without loss of generality, we assume that the random variable  $S \in \mathbb{R}^d$  has a probability density  $p(S)$ . For a given portfolio  $x$ , the probability of the loss not exceeding a threshold  $\alpha$  is given by the cumulative distribution function

$$\Psi(x, \alpha) \stackrel{\text{def}}{=} \int_{f(x, S) \leq \alpha} p(S) dS. \quad (1)$$

When the probability distribution for the loss has no jumps,  $\Psi(x, \alpha)$  is everywhere continuous with respect to  $\alpha$ .

VaR associated with a portfolio  $x$ , for a specified confidence level  $\beta$  and time horizon  $\bar{t}$ , is given by

$$\alpha_\beta(x) \stackrel{\text{def}}{=} \inf\{\alpha \in \mathbb{R} : \Psi(x, \alpha) \geq \beta\}. \quad (2)$$

Note that under the assumption that  $\Psi(x, \alpha)$  is everywhere continuous, there exists  $\alpha$  (possibly not unique) such that  $\Psi(x, \alpha) = \beta$ .

Define  $[f(x, S) - \alpha]^+$  as

$$[f(x, S) - \alpha]^+ \stackrel{\text{def}}{=} \begin{cases} f(x, S) - \alpha & \text{if } f(x, S) - \alpha > 0, \\ 0 & \text{otherwise.} \end{cases}$$

The risk measure CVaR,  $\phi_\beta(x)$ , is defined as (Pflug, 2000; Rockafellar and Uryasev, 2002)

$$\phi_\beta(x) \stackrel{\text{def}}{=} \inf_\alpha (\alpha + (1 - \beta)^{-1} \mathbb{E}([f(x, S) - \alpha]^+)).$$

When the loss distribution has no jumps, CVaR is the conditional expectation of the loss, given that the loss is  $\alpha_\beta(x)$  or greater, and is given by

$$\phi_\beta(x) = (1 - \beta)^{-1} \int_{f(x,S) \geq \alpha_\beta(x)} f(x, S) p(S) dS. \tag{3}$$

Define the augmented function

$$F_\beta(x, \alpha) \stackrel{\text{def}}{=} \alpha + (1 - \beta)^{-1} \int_{S \in \mathbb{R}^d} [f(x, S) - \alpha]^+ p(S) dS. \tag{4}$$

Under the assumption that the loss function  $f(\cdot, S)$  is convex and the loss distribution is continuous, it can be shown (Rockafellar and Uryasev, 2000) that function  $F_\beta(x, \alpha)$  is convex and continuously differentiable with respect to  $\alpha$  and  $\phi_\beta(x)$  is convex with respect to  $x$ . Moreover, minimizing CVaR over any  $x \in X$ , where  $X$  is a subset of  $\mathbb{R}^n$ , is equivalent to minimizing  $F_\beta(x, \alpha)$  over  $(x, \alpha) \in X \times \mathbb{R}$ , i.e.,

$$\min_{x \in X} \phi_\beta(x) \equiv \min_{(x, \alpha) \in X \times \mathbb{R}} F_\beta(x, \alpha). \tag{5}$$

If, in addition,  $X$  is a convex set, then the CVaR minimization problem

$$\min_{(x, \alpha) \in X \times \mathbb{R}} F_\beta(x, \alpha) \tag{6}$$

is a convex programming problem.

### 3. Minimizing risk for derivative portfolios

At a given time horizon  $\bar{t} > 0$ , assume that the underlying asset prices of the derivative instruments are  $S_t \in \mathbb{R}^d$ , the initial asset prices are  $S_0$ , and the function  $f(x, S)$  is the loss of a portfolio from a universe of  $n$  instruments. Assume that instrument values at time  $\bar{t}$  are  $\{V_1(S_{\bar{t}}, \bar{t}), \dots, V_n(S_{\bar{t}}, \bar{t})\}$ . For a portfolio selection problem and a given investment horizon  $\bar{t} > 0$ , the loss associated with the portfolio  $x$  is

$$f(x, S_{\bar{t}}) = -x^T (V^{\bar{t}} - V^0),$$

where for any time  $t$ ,  $V^t \stackrel{\text{def}}{=} [V_1(S_t, t), \dots, V_n(S_t, t)]$ . Note that  $f(x, S)$  is a linear function of  $x$  and it can be easily shown that, for any  $\rho > 0$ ,

$$\alpha_\beta(\rho \cdot x) = \rho \cdot \alpha_\beta(x), \quad \text{and} \quad \phi_\beta(\rho \cdot x) = \rho \cdot \phi_\beta(x).$$

Let  $\delta V \in \mathbb{R}^n$  denote the change in the instrument values over the time horizon  $\bar{t}$ , i.e.,  $\delta V = V^{\bar{t}} - V^0$ . Then the loss,  $f(x, S_{\bar{t}})$ , of the portfolio over the investment horizon  $\bar{t}$  is  $-(\delta V)^T x$ .

Without loss of generality, let  $x \in \mathbb{R}^n$  denote the ratio of the instrument holdings to the total initial investment wealth, i.e.,  $x_i$  is the number of units of the  $i$ th instrument holding per dollar investment. (VaR and CVaR of a portfolio with a budget  $\rho$  are simply  $\rho \cdot \alpha_\beta(x)$  and  $\rho \cdot \phi_\beta(x)$ , respectively, where  $\alpha_\beta(x)$  and  $\phi_\beta(x)$  are computed for a dollar's investment).

Assume for now that the only constraints on the optimal portfolio are the budget constraint

$$(V^0)^T x = 1$$

and the return constraint for the investment horizon  $\bar{t}$

$$(\bar{\delta V})^T x = \bar{r},$$

where  $\bar{r} \geq 0$  specifies the expected return of the portfolio over the time horizon  $\bar{t}$  and  $\bar{\delta V} \in \mathbb{R}^n$  is the expected gain for the instruments, i.e.,  $\bar{\delta V} = \mathbb{E}[(\delta V)]$ .

If  $X = \{x : (V^0)^T x = 1, (\bar{\delta V})^T x = \bar{r}\}$  is the set of feasible portfolios, we can write (6) explicitly as

$$\begin{aligned} \min_{(x,z)} & \left( \alpha + (1 - \beta)^{-1} \int_{S \in \mathbb{R}^d} [-(\delta V)^T x - \alpha]^+ p(S) dS \right) \\ \text{subject to} & \quad (V^0)^T x = 1 \quad \text{and} \quad (\bar{\delta V})^T x = \bar{r}. \end{aligned} \tag{7}$$

We assume that a stochastic model for changes of the underlying asset prices of all the instruments in a portfolio is given. In addition, we assume that there exist methods for computing the derivative values, such as Black–Scholes formulae, delta–gamma approximations, and Monte Carlo simulation.

The continuous CVaR optimization problem (7) is a convex nonlinear minimization problem with linear constraints. If the loss distribution function is continuous, the objective function is continuously differentiable. How well is this optimization problem posed for portfolios of derivatives?

To investigate this, let us consider the delta–gamma approximation of derivative values. For a short time horizon  $t > 0$ , a delta–gamma approximation can be a sufficiently accurate approximation to the derivative value and is often used in risk assessment. In general, the delta–gamma approximation describes the most significant component in the change of the derivative values and can thus provide insight into the nature of the solution. Thus, we assume for now that the change, for a given horizon  $\bar{t}$ , in instrument values is specified by the delta–gamma approximation: for instrument  $i$ ,

$$V_i^{\bar{t}} - V_i^0 = \left( \frac{\partial V_i^0}{\partial t} \right) \delta \bar{t} + \left( \frac{\partial V_i^0}{\partial S} \right)^T (\delta S) + \frac{1}{2} (\delta S)^T \Gamma_i (\delta S). \tag{8}$$

Here the vector  $(\delta S) \in \mathbb{R}^d$  denotes the change in the underlying values,  $\frac{\partial V_i^0}{\partial t}$  denotes the initial theta sensitivity of the  $i$ th instrument value to time,  $\frac{\partial V_i^0}{\partial S} \in \mathbb{R}^d$  denotes the initial delta sensitivity of the  $i$ th instrument with respect to the underlyings, and  $\Gamma_i \in \mathbb{R}^{d \times d}$  is the Hessian matrix denoting the initial gamma sensitivity of the  $i$ th instrument with respect to the underlyings, and  $\delta t$  is change in time.

Let  $\frac{\partial V^0}{\partial t}$  and  $\frac{\partial V^0}{\partial S}$  denote the initial sensitivities for all instruments in the investment universe:

$$\begin{aligned} \frac{\partial V^0}{\partial t} &\stackrel{\text{def}}{=} \left[ \frac{\partial V^0_1}{\partial t}, \dots, \frac{\partial V^0_n}{\partial t} \right] \in \mathbb{R}^n, \\ \frac{\partial V^0}{\partial S} &\stackrel{\text{def}}{=} \left[ \frac{\partial V^0_1}{\partial S}, \dots, \frac{\partial V^0_n}{\partial S} \right]^T \in \mathbb{R}^{n \times d}. \end{aligned}$$

Assume for now that each instrument depends on a single risky asset. If a derivative value depends on more than one risk factor, similar results can be obtained by accounting for the cross-partial derivatives; this analysis is presented in [Appendix A](#). In the case of a single risk factor, the only non-zero entries in the vector  $\frac{\partial V^0}{\partial S}$  and matrix  $\Gamma_i$  are entries  $i$  and  $(i,i)$ , respectively. Let

$$\Gamma \stackrel{\text{def}}{=} [\Gamma_1^{\text{diag}}, \dots, \Gamma_n^{\text{diag}}]^T \in \mathbb{R}^{n \times d},$$

where  $\Gamma_i^{\text{diag}}$  represents the diagonal of the matrix  $\Gamma_i$  as a column vector. Let  $(\delta S)^2$  be the vector with each entry of  $\delta S$  squared. If we set

$$A \stackrel{\text{def}}{=} \left[ \left( \frac{\partial V^0}{\partial t} \right), \left( \frac{\partial V^0}{\partial S} \right), \frac{1}{2} \Gamma \right] \in \mathbb{R}^{n \times (2d+1)}, \tag{9}$$

the loss in portfolio value is given by

$$f(x, S) = -x^T A \begin{bmatrix} \delta t \\ \delta S \\ (\delta S)^2 \end{bmatrix}. \tag{10}$$

If  $n > 2d + 1$ , there exists a non-zero  $z \in \mathbb{R}^n$  satisfying  $A^T z = 0$ . It is clear that, for any  $\theta$ ,

$$f(x, S) = f(x + \theta \cdot z, S), \quad \forall S.$$

Thus the portfolios  $x$  and  $(x + \theta \cdot z)$  have the same VaR and CVaR under the delta–gamma approximation. For a portfolio selection problem, if  $X = \{x : (V^0)^T x = 1, (\bar{\delta V})^T x = \bar{r}\}$  denotes the set of feasible portfolios corresponding to the budget and return constraints, we may deduce that if  $n > (2d + 3)$ , then the optimal CVaR and VaR portfolios for the selection problem defined by  $\min_{x \in X} \phi_\beta(x)$  and  $\min_{x \in X} \alpha_\beta(x)$ , for any  $0 < \beta < 1$ , lie in a linear subspace of dimension  $n - (2d + 3)$ . This implies that the VaR and CVaR derivative minimization problems, under these stated assumptions, are ill-posed and different computational methods may produce different optimal portfolios.

We note that not all derivative values can be simply expressed as  $V(S_t, t)$  where  $S_t$  represents risk factor values at time  $t$ . Asian options for example have a strong dependency on the history of the stock price. The analysis when such instruments are present is more complex but similar results may be obtained.

When the derivative values are computed through more accurate methods, such as analytic formulae or Monte Carlo simulation, the CVaR optimization problem

typically remains ill-posed in the sense that there are many portfolios that have similar risk values to that of the optimal portfolio and slight perturbation of the data can lead to significantly different optimal solutions.

Although much of the subsequent discussion is applicable to the VaR optimization problem, we will focus only on the CVaR optimization problem due to its computational tractability.

A continuous CVaR optimization problem (7) can be approximated using Monte Carlo simulation. Assume that  $\{(\delta V)_i\}_{i=1}^m$  are independent samples of  $\delta V$ , the change in the instrument values over the given horizon. Then the following is an approximation to the optimization problem (7):

$$\min_{(x,z) \in X \times \mathbb{R}} \left( \bar{F}_\beta(x, \alpha) \stackrel{\text{def}}{=} \alpha + \frac{1}{m(1-\beta)} \sum_{i=1}^m [-(\delta V)_i^\top x - \alpha]^+ \right). \quad (11)$$

When the subset  $X$  is specified by a finite set of linear constraints, (11) has an equivalent linear programming formulation which can be solved by standard methods for linear programming, see e.g., Rockafellar and Uryasev (2000) for an overview of modeling the CVaR problem (as well as other risk measures) as an LP. Here we use interior point method software MOSEK (Mosek, 2002).

Naturally, additional properties can be included in the CVaR optimization problem as constraints to alleviate the ill-posedness of the problem and produce a more desirable optimal portfolio. However, one needs to be careful to ensure that these constraints are meaningful and consistent in the sense that there exist feasible solutions. In addition, simply adding constraints may give a false sense of security; the optimization problem may remain ill-posed, as will be illustrated next.

The most natural constraints that one can add are simple bound constraints on the instrument holdings. The following example illustrates that this does not necessarily regularize the ill-posedness of the problem. In addition, it demonstrates significant consequences of the ill-posedness of the CVaR derivative portfolio optimization problem.

Assume, for example, that the feasible portfolios satisfy budget and return constraints as well as bound constraints on the instrument positions. Then the optimization problem (11) becomes

$$\begin{aligned} \min_{(x,z)} & \left( \alpha + \frac{1}{m(1-\beta)} \sum_{i=1}^m [-(\delta V)_i^\top x - \alpha]^+ \right) \\ \text{subject to} & \quad (V^0)^\top x = 1, \quad (\bar{\delta V})^\top x = \bar{r}, \quad \text{and } l \leq x \leq u. \end{aligned} \quad (12)$$

To illustrate the properties of the optimal portfolio from (12), we consider a universe of 196 instruments consisting of 12 vanilla calls, 12 vanilla puts, 12 binary calls, and 12 binary puts on each of the four correlated assets, and the four underlying assets themselves; here the derivative instruments are all European options. The initial asset prices, the covariance matrix of the annual returns, and the expected rates of return of the four assets are given in Table 7 in Appendix B. In this paper, we use these specifications for all our computational results, varying only the option types being

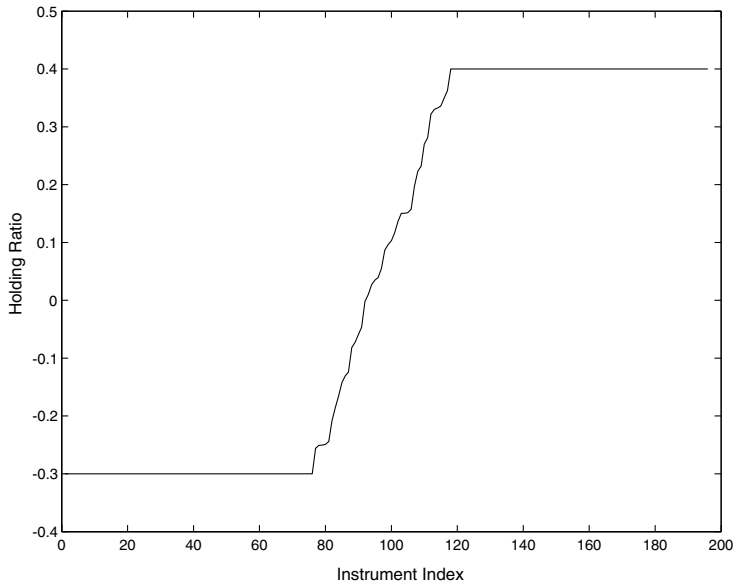


Fig. 1. (ordered) Holding ratios from (12) with no cost consideration:  $\beta = 0.95$ ,  $\bar{t} = 10$  days.

considered. The derivatives are priced using Black–Scholes type formulae, assuming that the underlying prices are log-normally distributed. For this example, we use  $m = 25,000$  Monte Carlo samples. The strike prices used for options on each asset are  $[0.8; 1.025; 1.25] \times S_0$  where  $S_0$  is the time 0 asset values (see Table 7 in Appendix B). The times to expiry are  $[2; 4; 6; 8] \times \bar{t}$ , where  $\bar{t} = 10$  days is the investment horizon (we assume that there are 250 trading days in a year). The options are all combinations of strikes and expiry times. The required portfolio return is twice the risk free interest rate over the investment horizon with the annual risk free interest rate equal to 5%. We use lower bounds of  $-0.3$  and upper bounds of  $0.4$  for this example. For an investment of \$100, no more than 30 units of each instrument can be shorted and no more than 40 units of each instrument can be bought. An (ordered) optimal portfolio holding ratio  $x^*$  from (12), computed using the interior point software MOSEK, is shown in Fig. 1.<sup>1</sup> It should be noted that the optimal portfolio computed using CPLEX was very similar to the one computed using MOSEK. In particular, the CPLEX optimal portfolio did not have any more instruments at their bounds than the MOSEK optimal portfolio.

Let us exclude the  $i$ th instrument from the optimal portfolio if  $|x_i^*| \leq 10^{-5}$ . We first observe that the optimal portfolio consists of all the instruments in the investment universe. In addition, about 77% of instrument holding ratios are equal to either their upper or lower limits. Such an optimal portfolio is undesirable in that it leads to large transaction as well as management costs. Moreover, any model error

<sup>1</sup> The instruments are ordered so as to illustrate the number of holding ratios that are significantly larger than zero, in particular, the number of holding ratios that are at their bounds.



will be magnified for a portfolio with extreme holdings; this is illustrated for the portfolio hedging problem in Alexander et al. (2003). Indeed, the optimization problem with the bound constraints remains ill-posed in the sense that there are many different portfolios with similar CVaR values; we will illustrate this next by providing a more desirable optimal CVaR formulation which produces more attractive portfolios with similar CVaR values.

In addition to the portfolio selection problem, another practically important derivative portfolio optimization problem is the portfolio hedging problem. In this context, for a given hedging horizon  $\bar{t}$ , one has an initial portfolio and an associated portfolio loss  $\Pi^0(S, \bar{t})$ . The goal is to decrease the risk of this portfolio by selecting an appropriate hedging portfolio from the available instruments  $\{V_1(S, t), \dots, V_n(S, t)\}$ . Thus the loss function for the hedging problem has the form

$$f(x, S) = \Pi^0(S, \bar{t}) - x^T(V(S, \bar{t}) - V(S_0, 0)).$$

While we devote our analysis of ill-posedness in this paper to the portfolio selection problem, similar analysis applies to the portfolio hedging problem where the goal is to hedge a given portfolio using more liquid derivatives with CVaR as the risk measure. Without mathematical analysis in the general setting and computational results for the smoothing technique, a shorter and simpler paper is written in Alexander et al. (2003) to illustrate the effect of the ill-posedness in the derivative portfolio hedging problem based on CVaR; hedging performances of the optimal portfolios under different cost considerations are also compared.

#### 4. CVaR optimization with cost

Given that the CVaR optimization problem for a portfolio of derivatives is ill-posed, additional meaningful criteria need to be considered for a derivative portfolio CVaR optimization problem in order to generate a stable solution. A natural meaningful consideration in portfolio investment or risk management is transaction and management cost. A portfolio, which, in addition to a small CVaR, incurs a small transaction and management cost, is certainly more attractive. We can regard the management cost as a function of the number of (non-zero holding) instruments in a portfolio. Unfortunately, it is difficult to include this explicitly into an optimization formulation since it is computationally challenging to solve the resulting mixed integer program. Our objective is to seek a portfolio which consists of a small number of instruments by minimizing a combination of CVaR and a suitable cost function without the need to solve a mixed integer programming problem.

Let us assume that the cost of holding an instrument is proportional to the magnitude of the instrument holdings. Then we seek a portfolio which has a minimum weighted combination of CVaR and the proportional cost:

$$\min_{x \in X} \left( \phi_\beta(x) + \sum_{i=1}^n c_i |x_i| \right), \quad (13)$$

where  $\phi_\beta(x)$  is as defined in (3). Here  $c \geq 0$  is the weighted cost, representing the cost as well as the tradeoff between minimizing CVaR and cost.

The weighted cost parameter  $c_i \geq 0$  can be interpreted as a measure of relative desirability to exclude the  $i$ th instrument from the optimal portfolio: if  $c_i$  is greater than some finite threshold value, and there exists a feasible portfolio with  $x_i = 0$ , then the optimal portfolio  $x^*$  for (13) is guaranteed to exclude the  $i$ th instrument, i.e.,  $x_i^* = 0$ . In this sense, we can regard our cost model as a model for management cost. This property of the cost model (13) is due to the fact that the objective function ( $\phi_\beta(x) + \sum_{i=1}^n c_i |x_i|$ ) is an exact penalty function of a constrained optimization problem. We refer interested readers to Fletcher (1981) for a more detailed discussion on the exact penalty function. Note that if one models the cost as  $\sum_{i=1}^n c_i x_i^2$  for example, the resulting optimal portfolio typically has few, if any at all, of its instruments with a small holding ratio  $|x_i^*|$  (e.g.,  $|x_i^*| \leq 10^{-5}$ ). For the quadratic penalty function, the constraint  $x_i^* = 0$  is only satisfied as the penalty parameter  $c_i$  tends to  $+\infty$ .

To solve (13), we can similarly consider the augmented function  $F_\beta(x, \alpha) + \sum_{i=1}^n c_i |x_i|$ . It is clear that  $F_\beta(x, \alpha) + \sum_{i=1}^n c_i |x_i|$  remains convex and continuously differentiable with respect to  $\alpha$  since  $\sum_{i=1}^n c_i |x_i|$  is convex and has no dependence on  $\alpha$ ; the analysis of Rockafellar and Uryasev (2000) applies. Moreover minimizing the sum of the weighted cost and CVaR of a portfolio  $x$  in any subset  $X$  of  $\mathbb{R}^n$  is equivalent to minimizing  $F_\beta(x, \alpha) + \sum_{i=1}^n c_i |x_i|$  over  $(x, \alpha) \in X \times \mathbb{R}$ , i.e.,

$$\min_{x \in X} \left( \phi_\beta(x) + \sum_{i=1}^n c_i |x_i| \right) \equiv \min_{(x, \alpha) \in X \times \mathbb{R}} \left( F_\beta(x, \alpha) + \sum_{i=1}^n c_i |x_i| \right).$$

In addition,  $F_\beta(x, \alpha) + \sum_{i=1}^n c_i |x_i|$  is convex with respect to  $(x, \alpha)$  and  $\phi_\beta(x) + \sum_{i=1}^n c_i |x_i|$  is convex with respect to  $x$  if the loss function  $f(x, S)$  is convex with respect to  $x$ . Moreover, if  $X$  is a convex set, the minimization problem

$$\min_{(x, \alpha) \in X \times \mathbb{R}} \left( F_\beta(x, \alpha) + \sum_{i=1}^n c_i |x_i| \right) \tag{14}$$

is a convex programming problem.

When (14) is approximated through Monte Carlo simulation, and  $X$  is specified by the budget and return constraints and bounds on the holding ratios  $x$ , the CVaR optimization problem with a proportional cost becomes a constrained piecewise linear minimization problem:

$$\begin{aligned} \min_{(x, \alpha)} & \left( \alpha + \frac{1}{m(1-\beta)} \sum_{i=1}^m [-(\delta V_i)^T x - \alpha]^+ + \sum_{j=1}^n c_j |x_j| \right) \\ \text{subject to} & \quad (V^0)^T x = 1, \quad (\overline{\delta V})^T x = \bar{r}, \quad \text{and } l \leq x \leq u. \end{aligned} \tag{15}$$

To illustrate the effect of the weighted cost parameter  $c$  on the optimal portfolio obtained from the CVaR cost model (14), we consider the weighted cost parameter  $c_i = \omega \cdot |\overline{\text{CVaR}}^0|$ ,  $1 \leq i \leq n$ , where  $\overline{\text{CVaR}}^0$  denotes the optimal CVaR from (15) with no cost consideration, for a dollar's investment. (Here we are implicitly assuming that the transaction costs of instruments are the same.)

Consider the same 196 instrument example in Section 3. We first recall that the optimal CVaR portfolio, under no cost consideration, contains all the 196 instruments. In addition 77% of holding ratios are at their bounds. Fig. 2 plots the optimal portfolio holding ratio  $x^*$ , for the same example, for the weighted cost  $c_i = \omega \cdot |\overline{\text{CVaR}}^0|$  where  $\omega = 0, 0.005, \text{ and } 0.01$ . We note that for  $\omega = 0.005$  and  $0.01$ , the optimal portfolios are preferable in the sense that they contain only 35.7% and 29.1% of the 196 instruments, respectively.

In order to analyze the impact of the cost consideration on risks, we consider the relative differences of VaR and CVaR under different weighted cost parameters with respect to that under cost consideration, i.e.,

$$\text{RelDifVaR}(\omega) \stackrel{\text{def}}{=} \left| \frac{\text{VaR}(\omega) - \text{VaR}^0}{\text{VaR}^0} \right|, \tag{16}$$

$$\text{RelDifCVaR}(\omega) \stackrel{\text{def}}{=} \left| \frac{\text{CVaR}(\omega) - \text{CVaR}^0}{\text{CVaR}^0} \right|, \tag{17}$$

where  $\text{CVaR}^0$  is the optimal CVaR value from (15) under no cost consideration. Throughout the paper, VaR and CVaR reported correspond to VaR and CVaR based on Monte Carlo simulations. Note that the loss distribution from simulations has jumps. Consider the loss associated with a portfolio  $x$  for  $m$  scenarios,  $(\text{loss})_1 \leq \dots \leq (\text{loss})_m$ , with each loss  $(\text{loss})_i$  having probability  $p_i$ . For a confidence level  $\beta$ , let  $i_\beta \leq m$  be the index such that

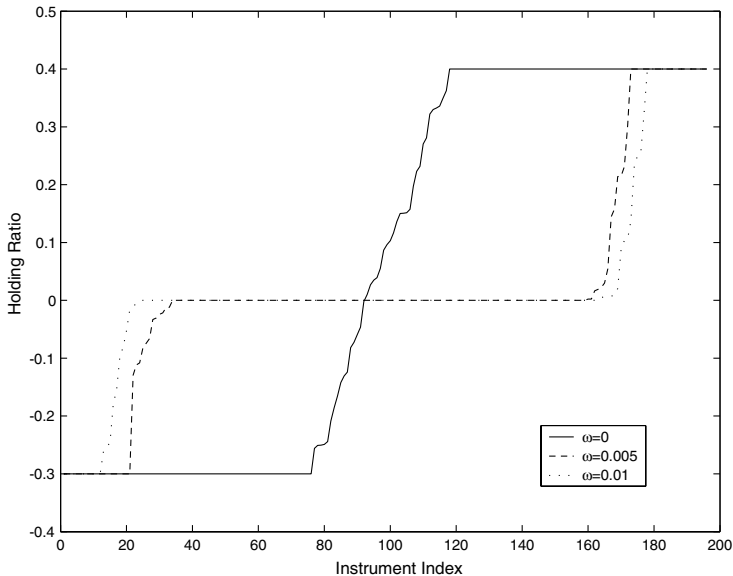


Fig. 2. Holding ratios with varying costs:  $\beta = 0.95, \bar{t} = 10$  days.

Table 1  
Effect of weighted cost parameters on the optimal CVaR portfolio for  $\bar{t} = 10$  days

$\beta$	$\omega$	VaR	CVaR	RelDifVaR	RelDifCVaR	# Instr
0.95	0.000	1.5795	1.6458	0.0000	0.0000	196
	0.005	1.6544	1.7251	0.0474	0.0482	70
	0.050	1.9933	2.0545	0.2620	0.2483	34
0.99	0.000	1.6891	1.7171	0.0000	0.0000	196
	0.005	1.7654	1.7910	0.0452	0.0431	72
	0.050	2.1271	2.1443	0.2593	0.2488	34

$$\sum_{i=1}^{i_\beta} p_i \geq \beta > \sum_{i=1}^{i_\beta-1} p_i. \tag{18}$$

Then VaR is given by  $\alpha_\beta(x) = (\text{loss})_{i_\beta}$  and CVaR equals

$$\phi_\beta(x) = \frac{1}{1-\beta} \left[ \left( \sum_{i=1}^{i_\beta} p_i - \beta \right) \alpha_\beta(x) + \sum_{i=i_\beta+1}^m p_i (\text{loss})_i \right]. \tag{19}$$

For a more detailed discussion of CVaR for scenario models, see [Rockafellar and Uryasev \(2002\)](#).

Table 1 tabulates relative risk differences for  $\beta = 0.95$  and  $\beta = 0.99$  with different weighted cost parameters for an investment horizon of 10 days. Computational results for longer maturities are similar and we do not report here. The investment universe consists of 196 instruments, see [Appendix B](#) for detailed specifications. VaR and CVaR reported here are for investment portfolios with an initial wealth of \$100. The results are for a single simulation problem and do not represent averages. Table 1 illustrates that, using the CVaR and cost optimization formulation (15), it is possible to obtain CVaR optimal portfolios with significantly fewer instruments but comparable risks. For example, for  $\bar{t} = 10$  days with  $\omega = 0.005$ , the optimal risks reflect an increase of less than 5% compared to that under no cost. Given the inevitable existence of model error as well as computational error due to, e.g., Monte Carlo approximation, a small difference in risk may be entirely acceptable. The number of instruments in the optimal portfolio, however, is less than 36% of the number of non-zero holdings under no cost, assuming a cutoff of  $10^{-5}$ . Not surprisingly, we observe that, as the cost parameter increases, the risk increases and the number of instruments in the optimal portfolio decreases.

### 5. Minimizing CVaR efficiently

The simulation CVaR optimization problem (15) is a piecewise linear minimization problem subject to linear constraints. As discussed previously, one way of computing a solution to (15) is to solve an equivalent linear programming problem:

$$\min_{(x,y,z,\alpha)} \left( \alpha + \frac{1}{m(1-\beta)} \sum_{i=1}^m y_i + \sum_{j=1}^n c_j z_j \right)$$

subject to

$$\begin{aligned} (V^0)^T x &= 1, \\ (\delta V)^T x &= \bar{r}, \\ y &\geq -Bx - \alpha e_m, \\ z - x &\geq 0, \quad z + x \geq 0, \\ l &\leq x \leq u, \quad y \geq 0, \end{aligned} \tag{20}$$

where the  $m$ -by- $n$  scenario loss matrix  $B$  is given by

$$B = [(\delta V)_1^T; (\delta V)_2^T; \dots; (\delta V)_m^T]$$

and  $e_m \in \mathbb{R}^m$  is the vector of all ones. This linear program has  $O(m + n)$  variables and  $O(m + n)$  constraints, where  $m$  is the number of Monte Carlo samples and  $n$  is the number of instruments. We assume that the loss  $(\delta V)$  is computed using computational methods such as analytic formulae and Monte Carlo simulation.

Linear programming is the simplest constrained optimization problem; there exists, for this class of problems, the most thorough theoretic analysis and reliable software. Although it is known that both CPLEX (a simplex type method) and MOSEK (an interior point method) are capable of solving very large linear programming problems in a short amount of time, the efficiency of both methods depends heavily on the *sparsity structures* of the problem. The linear programming problem arising from the CVaR optimization problem has a large dense block; the size of this dense block is determined by the number of scenarios and the number of instruments. We illustrate below that the computational cost for solving a CVaR problem via the linear programming approach quickly becomes prohibitive as the number of simulations and/or instruments become large.

Table 2 illustrates how the cpu time grows with the number of simulations and the number of instruments for the CVaR optimization problem (12). The comparison is made between CPLEX version 6.6 which implements a simplex method and the MOSEK Optimization Toolbox for MATLAB version 6 (for Solaris Sparc) which implements an interior point method for single problems. The problems are implemented in MATLAB version 6.1 and run on a Sun Sparc Ultra-2 machine.

Table 2  
CPU time for standard LP methods:  $\beta = 0.99$

# Scenarios	MOSEK (cpu sec)			CPLEX (cpu sec)		
	# of instruments being considered					
	8	48	200	8	48	200
10,000	11.07	61.96	1843.90	53.68	427.97	2120.84
25,000	30.02	162.13	14,744.64	351.44	2345.43	9907.99
50,000	43.62	642.24	–	1673.82	9296.98	–

Table 2 clearly illustrates that, using the standard linear programming software, the computational cost as well as the memory requirement quickly become prohibitive as the number of Monte Carlo samples and the number of instruments increase. For example, with 200 instruments and more than 25,000 simulations, a significant amount of the elapsed time is spent in swapping relevant data in and out of the cache memory. With 200 instruments and 50,000 scenarios, the elapsed time is significantly longer than that of the 48 instrument example, with the memory swapping dominating the elapsed time, and the entry is marked by “–” in the table.

As an alternative to the linear programming approach for the CVaR optimization problem, we investigate a computationally efficient method which directly exploits the property of the CVaR optimization problem; our ultimate objective is to be able to solve large scale CVaR portfolio problems.

We want to solve a portfolio CVaR optimization problem

$$\min_{(x,\alpha) \in X \times \mathbb{R}} \left( F_\beta(x, \alpha) + \sum_{i=1}^n c_i |x_i| \right)$$

through Monte Carlo simulation. We assume subsequently that the cumulative loss distribution function is continuous. The augmented CVaR function  $F_\beta(x, \alpha)$  is continuously differentiable under the assumption that the loss distribution has no jumps. The linear programming approach arises from approximating the continuously differentiable function  $F_\beta(x, \alpha)$  by the piecewise linear objective function

$$\bar{F}_\beta(x, \alpha) = \alpha + \frac{1}{m(1-\beta)} \sum_{i=1}^m [-(\delta V)_i^T x - \alpha]^+.$$

As the number of Monte Carlo simulations increases, the piecewise linear approximation  $\bar{F}_\beta(x, \alpha)$  approaches the continuously differentiable function  $F_\beta(x, \alpha)$ .

As an alternative to the piecewise linear approximation  $\bar{F}_\beta(x, \alpha)$ , we consider a continuously differentiable piecewise quadratic approximation  $\tilde{F}_\beta(x, \alpha)$  to the continuously differentiable function  $F_\beta(x, \alpha)$ . Let

$$\tilde{F}_\beta(x, \alpha) \stackrel{\text{def}}{=} \alpha + \frac{1}{m(1-\beta)} \sum_{i=1}^m \rho_\epsilon(-(\delta V)_i^T x - \alpha), \tag{21}$$

where  $\rho_\epsilon(z)$  is a continuously differentiable piecewise quadratic function which approximates the piecewise linear function  $\max(z, 0)$ : given a resolution parameter  $\epsilon > 0$ ,

$$\rho_\epsilon(z) \stackrel{\text{def}}{=} \begin{cases} z & \text{if } z \geq \epsilon, \\ \frac{z^2}{4\epsilon} + \frac{1}{2}z + \frac{1}{4}\epsilon & \text{if } -\epsilon \leq z \leq \epsilon, \\ 0 & \text{otherwise.} \end{cases} \tag{22}$$

To illustrate the smoothness of  $\bar{F}_\beta(x, \alpha)$  and  $\tilde{F}_\beta(x, \alpha)$ , let us consider the function  $g(\alpha) = E([S-\alpha]^+)$  assuming that  $S$  is a standard normal. Fig. 3 graphically illustrates the accuracy and smoothness of the approximations

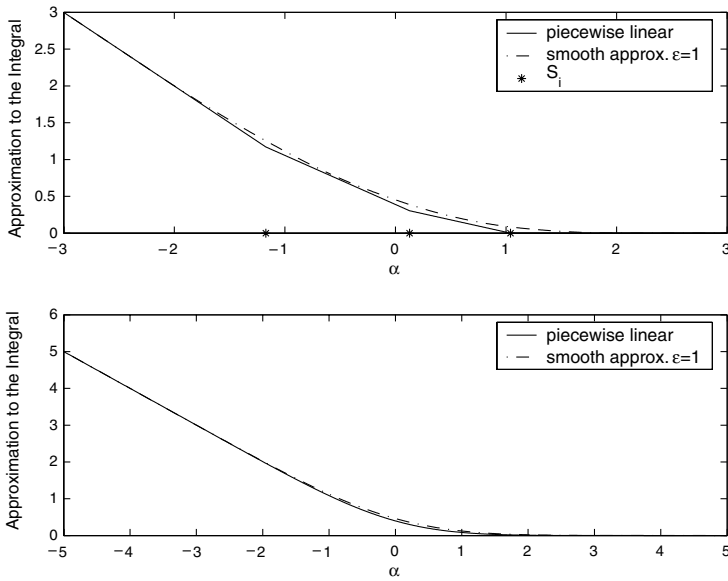


Fig. 3. A smooth approximation.

$$\frac{1}{m} \sum_{i=1}^m [S_i - \alpha]^+$$

and  $\frac{1}{m} \sum_{i=1}^m \rho_\epsilon(S_i - \alpha)$  as compared to  $g(\alpha)$ ; the top subplot is for  $m = 3$  (the asterisks on the  $x$ -axis represent  $S_i$ ) and the bottom subplot is for  $m = 10,000$ . It can be observed that, as the number of independent samples  $m$  increases, the difference between  $\frac{1}{m} \sum_{i=1}^m [S_i - \alpha]^+$  and  $\frac{1}{m} \sum_{i=1}^m \rho_\epsilon(S_i - \alpha)$  becomes smaller. In addition the function  $\frac{1}{m} \sum_{i=1}^m [S_i - \alpha]^+$  appears smoother.

Using  $\tilde{F}(x, \alpha)$  as a continuously differentiable approximation to  $F(x, \alpha)$ , we solve the following continuous piecewise quadratic convex programming problem

$$\begin{aligned} \min_{(x, \alpha)} & \left( \tilde{F}_\beta(x, \alpha) + \sum_{j=1}^n c_j |x_j| \right), \\ \text{subject to} & \quad (V^0)^T x = 1, \quad (\overline{\delta V})^T x = \bar{r}, \quad \text{and } l \leq x \leq u. \end{aligned} \tag{23}$$

Note that, for (20), each simulation introduces an additional variable (and constraint) in its equivalent linear program formulation when the piecewise linear function  $\bar{F}(x, \alpha)$  is used to approximate  $F(x, \alpha)$ . On the other hand, when the continuously differentiable function  $\tilde{F}(x, \alpha)$  is used to approximate  $F(x, \alpha)$ , problem (23) has  $(n + 1)$  independent variables and its equivalent nonlinear program formulation only has  $O(n)$  independent variables and constraints.

An optimization method for a convex nonlinear programming problem (23) typically generates an infinite sequence of approximations converging to a solution. At each iteration, however, it typically requires a function and a gradient evaluation and

$O(n^3)$  linear algebraic operations. The function/gradient evaluation costs  $O(mn)$ . If exact second-order derivatives are used, then the Hessian calculation in the worst case is  $O(n^2k)$  where  $k$  is the total number of simulations with  $|\text{tr}(\delta V)_i^T x - \alpha| \leq \epsilon$ . Given that CVaR optimization minimizes the tail loss with a typical confidence level of  $\beta \geq 0.9$ ,  $k$  is usually very small relative to  $m$  for most iterations.

Table 3 makes a comparison between the cpu times of the proposed smoothing formulation and the linear programming approach (interior point method software MOSEK is used here) for individual problems. We consider the derivative portfolio CVaR optimization problem whose investment universe consists of vanilla call and put options on the same four correlated assets described in Section 4 with the strikes and maturities described in Table 8 in Appendix B. The implementation of the smoothing method is based on an interior point method (Coleman and Li, 1996) for nonlinear minimization with bound constraints and is implemented in MATLAB v6.1. The comparison is made on a Sun Sparc Ultra-5\_10 machine. We observe that the smoothing method is much more efficient than the linear programming approach with up to a 1187% efficiency speedup. In addition, the 200 instruments and 50,000 simulations example can now be solved in less than 35 cpu minutes with the smooth formulation due to less memory requirement and better computational efficiency.

Next, we illustrate the accuracy and computational efficiency of this smoothing technique in greater detail. We consider here a different set of derivative portfolios on the same four correlated assets described in Section 4. The portfolios consist of an equal number of vanilla calls, vanilla puts, binary calls and binary puts on each asset. Once again the options are specified by all combinations of strikes ( $K_n \times S_0$ ) and expiry times ( $T_n \times \bar{t}$ ), where  $n$  is the number of instruments (description in Table 8 in Appendix B.) The investment horizon we use here is  $\bar{t} = 62.5$  days.

For various costs, Table 4 compares the CPU usage of MOSEK and our smoothing technique. We observe that the smoothing technique is more efficient compared to the linear programming method software MOSEK; the best cpu efficiency speedup is achieved with no cost consideration. However, with a larger parameter of  $\omega = 0.01$ , MOSEK requires less cpu than when  $\omega = 0.005$ , possibly due to the improved conditioning of the problem for a larger  $\omega$ .

For comparison, we consider the relative difference in risks  $Q_{VaR}$  and  $Q_{CVaR}$ , where  $Q_{VaR}$  is defined as

Table 3  
CPU times for MOSEK vs. smoothing:  $\beta = 0.99$

# scenarios	MOSEK (cpu sec)			Smoothing (cpu sec)		
	# of instruments being considered					
	8	48	200	8	48	200
10,000	6.47	42.04	4244.30	2.45	16.78	419.52
25,000	33.50	98.91	10,784.10	5.37	35.48	838.15
50,000	36.01	318.72	–	9.90	62.08	2080.16



Table 4

CPU times for varying costs and problem sizes,  $\beta = 0.95$ ,  $\bar{t} = 62.5$  days, and resolution parameter  $\epsilon = 0.005$

$m$	$n$	$\omega = 0$		$\omega = 0.005$		$\omega = 0.01$	
		MOSEK	SMTH	MOSEK	SMTH	MOSEK	SMTH
25,000	20	49.60	10.57	49.97	14.91	48.31	14.62
	100	826.95	92.58	885.72	267.14	687.96	177.41
	196	7484.89	875.29	4141.79	851.79	2258.44	1088.83
50,000	20	129.67	24.40	120.99	35.15	124.37	47.16
	100	2893.16	182.08	1242.03	449.98	1068.60	412.04
	196	–	1413.31	–	1672.05	–	1545.49

$$Q_{\text{VaR}} \stackrel{\text{def}}{=} \frac{\text{VaR}_s - \text{VaR}_m}{|\text{VaR}_m|}, \quad (24)$$

where  $\text{VaR}_s$  and  $\text{VaR}_m$  are VaR values computed by the smoothing technique and MOSEK, respectively. Note that we report CVaR as well as VaR since minimizing CVaR typically leads to a small VaR and VaR itself is an important risk measure. It is implicitly assumed that the number of scenarios and  $\omega$  are fixed in computing  $Q_{\text{VaR}}$ . The quotient  $Q_{\text{CVaR}}$  is defined in a similar manner.

Table 5 compare risks of the optimal portfolios computed by linear programming method software MOSEK and the smoothing method for different weighted cost parameters; we observe that the relative difference is less than 1.5% for this example.

For smoothing method, the continuously differentiable approximation  $\tilde{F}(x, \alpha)$  to  $F(x, \alpha)$  depends on the resolution parameter  $\epsilon$ . Typically this parameter is set to a value between 0.05 to 0.005; the resolution parameter value should be smaller for a larger number of simulations since this leads to better approximation. The resolution parameter of 0.005 typically leads to a negligible difference in optimal risks between the portfolios computed from the linear programming method and the smoothing method. Table 6 illustrates the effect of the resolution parameter  $\epsilon$ , on the cpu requirement and relative risk difference to that computed from MOSEK. We make a few interesting observations. Firstly, the risks computed from the smoothing method can be smaller than those computed by MOSEK; this suggests that the smooth

Table 5

Comparison of VaR/CVaR values computed by MOSEK and the proposed smoothing technique for  $\beta = 0.95$ ,  $\bar{t} = 62.5$  days, and resolution parameter  $\epsilon = 0.005$

$m$	$n$	$\omega = 0$		$\omega = 0.005$	
		$Q_{\text{VaR}}$ (%)	$Q_{\text{CVaR}}$ (%)	$Q_{\text{VaR}}$ (%)	$Q_{\text{CVaR}}$ (%)
25,000	20	-0.6701	0.2114	-0.7111	0.1984
	100	-0.7007	0.9433	-1.3946	0.2442
	196	-1.1442	1.4990	-0.6874	0.1637
50,000	20	-0.5598	-0.2362	-0.4370	0.0520
	100	-0.7407	0.3438	-0.8173	0.1572

Table 6

Comparison of VaR/CVaR values computed by MOSEK and the proposed smoothing technique for different resolution parameters  $\epsilon$ ,  $\omega = 0.01$ ,  $\beta = 0.95$ , and  $\bar{t} = 62.5$  days

$m$	$n$	$Q_{\text{VaR}}$ (%)	$Q_{\text{CVaR}}$ (%)	CPU time
$\epsilon = 0.005$				
25,000	20	-0.7348	0.1856	14.62
	100	-0.5261	0.8342	177.41
	196	-0.5941	1.7078	1088.83
50,000	20	-0.4684	0.0385	47.16
	100	-1.0684	-0.1008	412.04
	196	-	-	1545.49
$\epsilon = 0.0005$				
25,000	20	-0.0005	0.0000	33.61
	100	-0.0015	0.0001	238.13
	196	-0.0028	0.0007	647.71
50,000	20	-0.0009	-0.0002	48.34
	100	-0.0051	-0.0012	449.56
	196	-	-	1720.36

Table 7

Initial prices, annual expected rate of return, and covariance matrix

<i>Initial asset prices</i>			
100	50	30	100
<i>Annual expected rate of return</i>			
0.1091	0.0619	0.0279	0.0649
<i>Annual covariance matrix</i>			
0.2890	0.0690	0.0080	0.0690
0.0690	0.1160	0.0200	0.0610
0.0080	0.0200	0.0220	0.0130
0.0690	0.0610	0.0130	0.0790

approximation is an acceptable approximation to the augmented CVaR function, if not more preferable. Secondly, as  $\epsilon$  becomes smaller, the relative risk difference becomes smaller, as one might have expected.

## 6. Concluding remarks

In this paper, we analyze the well-posedness of the derivative portfolio risk minimization problem with CVaR and VaR as the choice of risk measures. We illustrate that this minimization problem is typically ill-posed for derivative portfolios. In particular, we have shown that, when the derivative values are computed through delta-gamma approximations, there typically are an infinite number of portfolios with the

Table 8  
Strike equals  $K_n \times S_0$  and expiries equal  $T_n \times \bar{t}$  for  $\bar{t} = 10$  days

<i>8 instrument example</i>					
$K_8$	1				
$T_8$	4				
<i>48 instrument example</i>					
$K_{48}$	0.8	1	1.25		
$T_{48}$	2	4			
<i>200 instrument example</i>					
$K_{200}$	0.8	0.9125	1.025	1.1375	1.25
$T_{200}$	2	3.5	5	6.5	8
<i>20 instrument example (underlying assets included)</i>					
$K_{20}$	1				
$T_{20}$	4				
<i>100 instrument example (underlying assets included)</i>					
$K_{100}$	0.9	1	1.1		
$T_{100}$	3	6			
<i>196 instrument example (underlying assets included)</i>					
$K_{196}$	0.8	1.025	1.25		
$T_{196}$	2	4	6	8	

same VaR and CVaR. Thus, the derivative portfolio selection problem of minimizing risk subject to a specified return typically has an infinite number of solutions when using delta–gamma approximations. When the derivative values are computed using more accurate methods such as Black–Scholes formulae and Monte Carlo techniques, the optimal CVaR or VaR problem is typically ill-posed.

We illustrate that one may not be able to remove the ill-posedness of the CVaR/VaR optimization problem by simply adding constraints. When simple bound constraints are imposed on the instrument holdings, the optimal CVaR portfolio typically has a large number of non-zero instrument holdings (mostly at their bounds). This type of optimal portfolio may not be desirable and can be problematic since it may entail large management and transaction costs, and it tends to magnify modeling error.

We propose the inclusion of a weighted cost consideration in the CVaR optimization problem. We model the cost as proportional to the magnitude of instrument holding; this cost model is capable of controlling transaction cost as well as management cost. We illustrate that minimizing CVaR together with this cost model leads to more desirable portfolios with significantly smaller transaction costs, fewer non-zero instrument holdings, and comparable CVaR (and VaR) measures.

We propose a computationally efficient method for solving a simulation based CVaR optimization problem by exploiting the fact that the objective function in the CVaR optimization problem approaches a continuously differentiable function as the number of Monte Carlo samples increases to infinity. With a preliminary implementation of the proposed method in MATLAB, a comparison is made with

the standard linear programming approach. We illustrate that solving a continuously differentiable piecewise quadratic approximation to the CVaR optimization problem is much more efficient, producing an optimal CVaR, for appropriately chosen resolution parameters, very close to that obtained without cost consideration. Furthermore, it is more suitable for solving large scale CVaR portfolio optimization problems.

Although we have focused, in this paper, on the optimal derivative portfolio investment problem for both theoretical analysis and computational illustrations, similar analysis and computational results on the effectiveness of the CVaR and cost minimization formulation are presented in Alexander et al. (2003) for derivative portfolio hedging problems.

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**Appendix A. CVaR and delta–gamma approximation**

In Section 3 we assumed that each instrument depended on a single risk factor. In general, an instrument value may depend on more than one risk factor. Assuming there are  $d$  risk factors, in the analysis that follows we discuss the conditioning of minimizing derivative portfolio VaR/CVaR for this general case. We first need to set up a matrix of second-order sensitivities of derivatives to the underlying risk factors. Assume that for each  $h$ ,  $1 \leq h \leq n$ ,  $V_h$  depends on  $r_h$  risk factors, where  $r_h \leq d$ . Then, the second-order sensitivity matrix  $\Gamma_h$ , corresponding to  $V_h$ , has at most  $r_h^2$  non-zero entries. Further, since  $\Gamma_h$  is symmetric, it has at most  $\frac{r_h(r_h+1)}{2}$  distinct non-zero entries. Assuming an ordering on the risk factors, construct the sets  $R_h = \{(i,j): V_h \text{ is dependent on risk factors } i \text{ and } j, \text{ i.e., } \Gamma_h(i,j) \neq 0, i \leq j\}$  for  $h = 1, \dots, n$ . Then

$$(\delta S)^T \Gamma_h (\delta S) = \sum_{(i,i) \in R_h} \Gamma_h(i,i) \delta S_i^2 + \sum_{(i,j) \in R_h, i \neq j} 2\Gamma_h(i,j) \delta S_i \delta S_j.$$

Consider  $\hat{R} = \bigcup_{h=1}^n R_h$ . We also set an order on  $\hat{R}$  as follows: for  $(i_1, j_1), (i_2, j_2) \in \hat{R}$ ,

$$\begin{aligned} (i_1, j_1) < (i_2, j_2) & \text{ if } i_1 < i_2, \\ (i_1, j_1) < (i_2, j_2) & \text{ if } i_1 = i_2 \text{ and } j_1 < j_2. \end{aligned}$$

Now we construct the second-order sensitivity matrix  $\Gamma \in \mathbb{R}^{\hat{d} \times \hat{d}}$  for all instruments under consideration, where  $\hat{d}$  is the cardinality of  $\hat{R}$ . Row  $i$  of  $\Gamma$  corresponds to the  $i^{th}$  smallest element of  $\hat{R}$  which we shall refer to as  $\hat{R}_i$ . Now

$$\Gamma(l, h) \stackrel{\text{def}}{=} \begin{cases} 0 & \text{if } \widehat{R}_l \notin R_h, \\ \frac{\partial^2 V_h^0}{\partial S_i^2}, \text{ where } \widehat{R}_l = (i, i) & \text{if } \widehat{R}_l \in R_h, \\ 2 \frac{\partial^2 V_h^0}{\partial S_i \partial S_j}, \text{ where } \widehat{R}_l = (i, j) & \text{if } \widehat{R}_l \in R_h, i \neq j. \end{cases} \quad (25)$$

We also construct the vector  $\widehat{\delta S}^2 \in \mathbb{R}^{\hat{d}}$  which represents the required corresponding second-order change in risk factors,  $(\widehat{\delta S}^2)_i \stackrel{\text{def}}{=} \delta S_i \delta S_j$ , where  $\widehat{R}_l = (i, j)$ .

**Theorem A.1.** Assume that a portfolio is formed from instruments  $\{V_1, \dots, V_n\}$  and the underlying risk factors of  $\{V_1, \dots, V_n\}$  are  $\{S_1, \dots, S_d\}$ . For a fixed investment horizon  $\bar{t} > 0$ ,  $1 \leq i \leq n$ , assume that

$$V_i^{\bar{t}} - V_i^0 = \left( \frac{\partial V_i^0}{\partial t} \right) \delta \bar{t} + \left( \frac{\partial V_i^0}{\partial S} \right)^T (\delta S) + \frac{1}{2} (\delta S)^T \Gamma_i (\delta S).$$

Then the following is true:

1. If  $n > \hat{d} + d + 1$ , where  $\hat{d}$  is the cardinality of  $\widehat{R}$ , and there exists a portfolio whose  $VaR = VaR^*$  and  $CVaR = CVaR^*$ , where  $VaR^*$  and  $CVaR^*$  are the minimal VaR and CVaR, then there are an infinite number of optimal portfolios that have the minimal VaR and CVaR.
2. If  $X = \{x : (V^0)^T x = 1, (\overline{\delta V})^T x = \bar{r}\}$  and  $n > \hat{d} + d + 3$ , then the optimal CVaR and VaR portfolios defined by  $\min_{x \in X} \phi_\beta(x)$  and  $\min_{x \in X} \alpha_\beta(x)$ , for any  $0 < \beta < 1$ , lie in a linear subspace of dimension  $n - (\hat{d} + d + 3)$ .

**Proof.** The results in the theorem hold due to the fact that there are an infinite number of portfolios with the same VaR and CVaR under the assumed assumption.

We first note that if portfolios  $x^{(1)}, x^{(2)} \in \mathbb{R}^n$  satisfy

$$f(x^{(1)}, S) \equiv f(x^{(2)}, S), \quad \text{for all possible } S,$$

then these two portfolios have the same VaR value and the same CVaR value.

Let

$$A \stackrel{\text{def}}{=} \left[ \left( \frac{\partial V^0}{\partial t} \right), \left( \frac{\partial V^0}{\partial S} \right), \frac{1}{2} \Gamma \right] \in \mathbb{R}^{n \times (\hat{d} + d + 1)}. \quad (26)$$

The proof is straight forward from the observation that

$$\begin{aligned} f(x, S) &= -x^T \left( \left( \frac{\partial V^0}{\partial t} \right) \delta \bar{t} + \left( \frac{\partial V^0}{\partial S} \right) (\delta S) + \frac{1}{2} \Gamma (\widehat{\delta S}^2) \right), \\ &= -x^T A \begin{bmatrix} \delta \bar{t} \\ \delta S \\ (\widehat{\delta S}^2) \end{bmatrix}. \end{aligned}$$

Table 9

Strike equals  $K_n \times S_0$  and expiries equal  $T_n \times \bar{t}$  for  $\bar{t} = 62.5$  days

20 instrument example (underlying assets included)				
$K_{20}$	1			
$T_{20}$	2			
100 instrument example (underlying assets included)				
$K_{100}$	0.9	1	1.1	
$T_{100}$	2	4		
196 instrument example (underlying assets included)				
$K_{196}$	0.8	1.025	1.25	
$T_{196}$	1.5	2	3	4

If  $n > \hat{d} + d + 1$ , there exists a non-zero  $z \in \mathbb{R}^n$  satisfying

$$A^T z = 0.$$

Then  $f(x + \theta z, S) = f(x, S)$  for any  $S$  and  $\theta$ .

For the second result, similarly there exists a non-zero  $z \in \mathbb{R}^n$  which lies in the null space of  $[A, V^0, \bar{\delta}V]^T$  (this null space has dimension  $n - (\hat{d} + d + 3)$ ).

Note that in the worst case,  $\hat{d} = \frac{d(d+1)}{2}$ . This completes our proof.  $\square$

## Appendix B. Data specifications

Table 7 describes the initial underlying prices, the expected annual rates of return, and the covariance matrix of the rate of return, respectively.

Tables 8 and 9 describe the strike prices and the expiry of various portfolios used in the paper for different investment horizons; they are classified according to the total number of instruments in the portfolio universe.

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